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Preprint 2004/001

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WWW: http://www.mathematik/uni-stuttgart.de/preprints

ISSN 1613-8309

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Strong Laws of Large Numbers by Elementary Tauberian Arguments

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Abstract. For Kolmogorov's strong law of large numbers an alternative short proof is given which weakens Etemadi's condition of pairwise independence. The argument uses the known – and elementary – equivalence of (Cesàro) C_1 - and C_2 -summability for one-sided bounded sequences. Also other strong laws of large numbers are established, partially via Borel summability.

2000 Mathematics Subject Classification: 60 F 15

Key words: strong law of large numbers, Tauberian theorem, C_1 - and C_2 -summability, Borel summability, covariance, maximal-correlation coefficient

1 INTRODUCTION

A sequence $(X_n)_{n \in \mathbb{N}}$ of real random variables having finite expectations EX_n obeys a socalled strong law of large numbers, if

$$\frac{1}{n}\sum_{k=1}^{n} (X_k - EX_k) \to 0 \text{ almost surely (a.s.)},$$
(1)

i.e.,

$$C_1 - \lim(X_n - EX_n) = 0 \text{ a.s.}$$

For the classical strong law of large numbers of Kolmogorov, which concerns independent identically distributed (integrable) real random variables, Etemadi [6] gave an elementary proof together with a generalization to pairwise independence. He used Chebyshev's inequality and the Borel-Cantelli lemma in the context of suitable subsequences of partial sums. In this paper we give an alternative short proof of the result, under a weakened independence assumption (Theorem 2), using the monotone convergence theorem in series form and the well-known equivalence of (Cesàro) C_1 - and C_2 -summability for one-sided bounded sequences which can be proved in an elementary way. The core of the argument leads to a general theorem which generalizes several known results on strong laws of large numbers for square integrable random variables (Theorem 1). A corresponding argument using C_1 and Borel summability yields a strong law of large numbers in a context slightly more general than second order stationarity (Theorem 3) and further under the classical Cramér-Leadbetter condition. A comparison with related results is given.

2 STRONG LAWS OF LARGE NUMBERS

First we formulate a rather general result.

Theorem 1 Let $(X_n)_{n \in \mathbb{N}}$ be a sequence of square integrable real random variables satisfying

$$\inf(X_n - EX_n) > -\infty \quad (a.s.) \tag{2}$$
$$(e.g., X_n \ge 0, \sup EX_n < \infty),$$
$$\sum_{n=1}^{\infty} \frac{\operatorname{Var}(X_1 + \ldots + X_n)}{n^3} < \infty. \tag{3}$$

Then (X_n) obeys a strong law of large numbers, i.e., (1) holds.

Remark 1 a) With covariance $\Gamma(i, j) := E\left[(X_i - EX_i)(X_j - EX_j)\right], i, j \in \mathbb{N}$, one has

$$\operatorname{Var}(X_1 + \ldots + X_n) = \sum_{i,j \in \{1,\ldots,n\}} \Gamma(i,j).$$
 (4)

Thus Theorem 1 generalizes the discrete time version of the almost sure stability theorem in Loève [21], section 37.7, A, in so far as (2) and (3) weaken the assumptions

$$|X_n| \le c < \infty, \quad \sum \frac{1}{n^3} \sum_{i,j \in \{1,\dots,n\}} |\Gamma(i,j)| < \infty$$

there. A sufficient condition for (3) together with (4) is

$$\sum_{i,j\in\mathbb{N}}\frac{|\Gamma(i,j)|}{(i+j)^2} < \infty.$$

For, with

$$d_n := \sum_{i,j \in \{1,\dots,n\}} |\Gamma(i,j)|$$

and by partial summation,

$$\sum_{n=1}^{\infty} \frac{d_n}{n^3} = d_1 \sum_{k=1}^{\infty} \frac{1}{k^3} + \sum_{n=1}^{\infty} \left(\sum_{k=n+1}^{\infty} \frac{1}{k^3} \right) (d_{n+1} - d_n)$$

$$\leq 2|\Gamma(1,1)| + \sum_{n=1}^{\infty} \frac{1}{n^2} \sum_{j=1}^{n+1} |\Gamma(n+1,j)|$$

$$\leq 16 \sum_{n,j \in \mathbb{N}} \frac{|\Gamma(n,j)|}{(n+j)^2} < \infty.$$

b) Theorem 1 also generalizes Theorem 1 of Etemadi [7], mainly in so far as (3) weakens the assumptions

$$EX_i X_j \le EX_i EX_j, \ j > i,$$

$$\sum \frac{\operatorname{Var}(X_n)}{n^2} < \infty$$
(5)

there. For these two conditions imply (3) via

$$\sum_{n=1}^{\infty} \cdot \frac{1}{n^3} \operatorname{Var}\left(\sum_{k=1}^n X_k\right) \leq \sum_{n=1}^{\infty} \frac{1}{n^3} \left(\sum_{k=1}^n \operatorname{Var}(X_k)\right)$$
$$= \sum_{k=1}^{\infty} \left(\sum_{n=k}^{\infty} \frac{1}{n^3}\right) \operatorname{Var}(X_k) < \infty.$$

c) Note that the Rademacher-Menchoff theorem (Rademacher [25], Menchoff [22]; see Révész [26], § 3.2, and, in a generalization, Loève [21], section 36, and Stout [29], Theorem 3.7.2) states a.s. convergence of $\sum (X_n - EX_n)/n$ and thus (1) for a sequence $(X_n)_{n \in \mathbb{N}}$ of square integrable pairwise uncorrelated real random variables under the condition

$$\sum_{n=1}^{\infty} \frac{\operatorname{Var}(X_n)}{n^2} (\log n)^2 < \infty.$$

The proof of Theorem 1 is based on the following deterministic Lemma 1, which is a classic Tauberian theorem and can be proved in an elementary way (see section 3).

Lemma 1 Let the sequence $(c_n)_{n \in \mathbb{N}}$ of real numbers be bounded from below. If

$$\sum_{n=1}^{\infty} \frac{(c_1 + \dots + c_n)^2}{n^3} < \infty$$
(6)

or only

$$\frac{1}{n^2} \sum_{k=1}^n (c_1 + \ldots + c_k) \to 0,$$
(7)

then

$$\frac{1}{n}\sum_{i=1}^{n}c_{i} \to 0.$$
(8)

The following theorem can be easily deduced from Theorem 1, by usual truncation and use of the well-known Lemma 2. It contains Etemadi's [6] as well as Cohn's [4] generalizations of the classical strong law of large numbers of Kolmogorov [18] (Loève [21], section 17). Its formulation uses the maximal-correlation coefficient $\kappa = \kappa(X, Y)$ (Hirschfeld [13], Gebelein [10]) defined for a pair (X, Y) of random variables by

$$\kappa(X,Y): = \sup\left\{\frac{E(\tilde{X} - E\tilde{X})(\tilde{Y} - E\tilde{Y})}{\operatorname{Var}(\tilde{X})^{1/2}\operatorname{Var}(\tilde{Y})^{1/2}}; \quad \tilde{X} \in \mathcal{L}_{2}(\mathcal{F}(X)), \quad \tilde{Y} \in \mathcal{L}_{2}(\mathcal{F}(Y))\right\}$$
$$= \sup\left\{\frac{E\tilde{X}\tilde{Y} - E\tilde{X}E\tilde{Y}}{(E\tilde{X}^{2})^{1/2}(E\tilde{Y}^{2})^{1/2}}; \quad \tilde{X} \in \mathcal{L}_{2}(\mathcal{F}(X)), \quad \tilde{Y} \in \mathcal{L}_{2}(\mathcal{F}(Y))\right\}$$

where 0/0 := 0, $\mathcal{F}(X)$ denotes the σ -algebra generated by X and $\tilde{X} \in \mathcal{L}_2(\mathcal{F}(X))$ denotes an $\mathcal{F}(X)$ - \mathcal{B} -measurable square integrable real random variable (compare also Kolmogorov and Rozanov [19]). $\kappa(X, Y) = 0$ means independence of X and Y.

Theorem 2 Let $(X_n)_{n \in \mathbb{N}}$ be a sequence of identically distributed integrable real random variables. If $\kappa_{i,j} := \kappa(X_i, X_j)$, satisfies

$$\sum_{n} \sup_{m} \kappa_{m,m+n} < \infty \tag{9}$$

or only

$$\sup_{j \in \{2,3,\dots\}} \sum_{i=1}^{j-1} \kappa(j-i,j) < \infty,$$
(10)

then (1) holds.

Remark 2 a) For $\kappa_{i,j} = 0$ $(i, j \in \mathbb{N})$, i.e., pairwise independence of the $X'_n s$, Theorem 2 yields Etemadi's ([6], [7]) generalization of Kolmogorov's strong law of large numbers from independence to pairwise independence.

b) For a pair (X, Y) of random variables let

$$\phi(X,Y) := \sup\left\{\frac{|P(A\cap B) - P(A)P(B)|}{P(A)}; A \in \mathcal{F}(X), B \in \mathcal{F}(Y)\right\},\\\lambda(X,Y) := \sup\left\{\frac{|P(A\cap B) - P(A)P(B)|}{P(A)^{1/2}P(B)^{1/2}}; A \in \mathcal{F}(X), B \in \mathcal{F}(Y)\right\}.$$

Then

$$\kappa(X,Y) \le 2\phi(X,Y)^{1/2}$$

(Ibragimov [14], p. 351) and

$$\lambda(X,Y) \le \kappa(X,Y) \le 3000\lambda(X,Y)(1 - \log\lambda(X,Y))$$

(Bradley [2], p. 168, and Bradley and Bric [3], p. 337). Setting

$$\phi_n^* := \sup_m \phi(X_m, X_{m+n}),$$

$$\lambda_n^* := \sup_m \lambda(X_m, X_{m+n}),$$

$$\kappa_n^* := \sup_m \kappa(X_m, X_{m+n}) (= \sup_m \kappa_{m,m+n})$$

for a sequence $(X_n)_{n \in \mathbb{N}}$ of random variables, we immediately obtain that

$$\sum \phi_n^{*1/2} < \infty \text{ or } \sum \lambda_n^* |\log \lambda_n^*| < \infty$$

implies $\sum \kappa_n^* < \infty$, i.e., (9), which implies condition (10). c) For a sequence $(X_n)_{n\in\mathbb{N}}$ of random variables let $\mathcal{F}_1^m = \mathcal{F}(X_1, \ldots, X_m)$, $\mathcal{F}_{m+n} = \mathcal{F}(X_{m+n}, X_{m+n+1}, \ldots)$ be the σ -algebras generated by X_1, \ldots, X_m and $X_{m+n}, X_{m+n+1}, \ldots)$ respectively, and let

$$\phi(n) := \sup_{m} \sup_{A \in \mathcal{F}_1^m, B \in \mathcal{F}_{m+n}} \frac{|P(A \cap B) - P(A)P(B)|}{P(A)}$$

 $\phi(n) \to 0 \ (n \to \infty)$ means that $(X_n)_{n \in \mathbb{N}}$ is ϕ -mixing. Cohn [4] showed that for a sequence $(X_n)_{n \in \mathbb{N}}$ of identically distributed integrable real random variables (1) holds if $\sum \phi_n^{1/2} < \infty$ (see also Iosifescu and Theodorescu [15], p. 19). This result is a consequence of Theorem 2, because $\sum \phi_n^{1/2} < \infty$ implies $\sum \phi_n^{*1/2} < \infty$ and thus (9) by b).

Lemma 2 Let X be an integrable nonnegative random variable and set

$$X^{[n]} := XI_{[X \le n]}, \quad n \in \mathbb{N}$$

(I denoting an indicator function). Then

$$\sum_{n=1}^{\infty} \frac{E(X^{[n]^2})}{n^2} < \infty.$$

Remark 3 Theorem 1, Lemma 1, and Remark 1 can be generalized from arithmetic means to weighted means. Let $\alpha_n \in [0, 1), n \in \mathbb{N}$, satisfy $\alpha_n \to 0$, $\sum \alpha_n = \infty$, and set

$$\beta_n := \frac{1}{(1 - \alpha_n) \dots (1 - \alpha_n)}, \quad \gamma_n := \alpha_n \beta_n,$$

thus $\beta_n = 1 + \gamma_1 + \ldots + \gamma_n \uparrow \infty$. The special case $\alpha_n = 1/(n+1)$ leads to $\beta_n = n+1, \gamma_n = 1$. Replace assumption (3) in Theorem 1 by

$$\sum_{n=1}^{\infty} \frac{\gamma_n \operatorname{Var}(\gamma_1 X_1 + \ldots + \gamma_n X_n)}{\beta_n^3} < \infty,$$

assumptions (6) and (7) in Lemma 1 by

$$\sum_{n=1}^{\infty} \frac{\gamma_n (\gamma_1 c_1 + \ldots + \gamma_n c_n)^2}{\beta_n^3} < \infty$$

and

$$\frac{1}{\beta_n} \sum_{k=1}^n \gamma_k \frac{1}{\beta_k} \left(\sum_{j=1}^k \gamma_j c_j \right) \to 0,$$

respectively, and assumption (5) in Remark 1b by

$$\sum_{n=1}^{\infty} \alpha_n^2 \operatorname{Var}(X_n) < \infty.$$

Then the assertion (1) in Theorem 1, also in context of Remark 1 b, has to be replaced by

$$\frac{1}{\beta_n} \sum_{k=1}^n \gamma_k (X_k - EX_k) \to 0 \text{ a.s.}, \tag{11}$$

and the assertion (8) in Lemma 1 has to be replaced by

$$\frac{1}{\beta_n}\sum_{k=1}^n \gamma_k c_k \to 0$$

(as to the latter, compare Karamata [17], with review in Zentralblatt Math. 19, pp. 341, 342). Relation (11) is of interest in stochastic approximation (see, e.g., Ljung, Pflug, and Walk [20], I.1, I.2, with further references).

Condition (2) of one-sided boundedness in Theorem 1 can be avoided by sharpening now the condition on Γ appearing in (3) together with (4). The proof uses another Tauberian argument.

Proposition 1 Let $(X_n)_{n \in \mathbb{N}}$ be a sequence of square integrable real random variables satisfying

$$\sum_{k=0}^{\infty} \sum_{j=k}^{\infty} e^{-k^2/(6j)} j^{-3/2} |\Gamma(j+k,j)| < \infty,$$
(12)

where $\Gamma(i, j)$ is defined by (4). Then (1) holds.

The following theorem is a consequence of Proposition 1 and comprehends the case of second order stationarity where $\Gamma(i, j)$ only depends on the difference i - j.

Theorem 3 Let $(X_n)_{n \in \mathbb{N}}$ be a sequence of square integrable real random variables satisfying

$$|\Gamma(i,j)| \le r(|i-j|), i, j \in \mathbb{N}$$
(13)

with

$$\sum_{k=1}^{\infty} \frac{r(k)}{k} < \infty, \tag{14}$$

e.g., with

$$r(k) = O\left(\frac{1}{\log k (\log \log k)^{1+\delta}}\right)$$
(15)

for some $\delta > 0$. Then (1) holds.

Remark 4 In the case of second order stationarity with autocovariance function $R(k) := \Gamma(j+k,k)$, via the spectral measure of (X_n) , Gaposhkin [9] established (1) under the (in some sense) weakest possible condition that

$$\sum_{k=2}^{\infty} \frac{R(k)}{k \log k} \log \log k \text{ converges},$$

where R(k) also may be replaced by $\overline{R}(k) := (R(1) + \ldots + R(k))/k$. It is well-known and can easily be proved via Lemma 1 that for a second order stationary process (X_n) the condition $\overline{R}(n) \to 0$ $(n \to \infty)$ is necessary and sufficient for

$$E\left|\frac{1}{n}\sum_{k=1}^{n}(X_{k}-EX_{k})\right|^{2}\to 0 \ (n\to\infty).$$

Remark 5 Noticing

$$\sum_{i,j\in\{m+1,\dots,n\}} |\Gamma(i,j)| + \sum_{i,j\in\{n+1,\dots,N\}} |\Gamma(i,j)| \le \sum_{i,j\in\{m+1,\dots,N\}} |\Gamma(i,j)|$$

for m < n < N, from Serfling [28] (see also Stout [29], Theorem 3.7.3) one obtains that the condition

$$\sum_{i,j\in\{1,\dots,n\}} |\Gamma(i,j)| \le cn^2 (\log n \log \log n)^{-2}, n \in \mathbb{N}$$
(16)

for some $c \in \mathbb{R}_+$, is sufficient for (1). $|\Gamma|$ may be replaced by the positive part Γ_+ . In the special case (13), the conditions

$$\sum_{j=1}^{k} |r(j)| = O(k(\log k \log \log k)^{-2})$$

(compare Stout [29], Theorem 3.7.4 with proof) and

$$r(k) = O((\log k \log \log k)^{-2})$$

yielding (16), are stronger than (14) and (15), respectively.

The proof of Proposition 1 is based on the deterministic Lemma 4 below. This lemma will be proved by Lemma 3, which is a consequence of a classic and deep Tauberian theorem on Borel summability, but can be shown in an elementary way (see section 3).

Lemma 3 Let the sequence $(d_n)_{n \in \mathbb{N}_0}$ of real numbers satisfy

$$\sum_{n=1}^{\infty} n^{1/2} (d_n - d_{n-1})^2 < \infty.$$
(17)

If

$$e^{-\lambda} \sum_{j=0}^{\infty} \frac{1}{j!} \lambda^j d_j \to 0 \ (\lambda \to \infty),$$
 (18)

then

$$d_n \to 0 \ (n \to \infty). \tag{19}$$

Relation (18) denotes the so-called Borel summability of the series $d_0 + \sum_{k=0}^{\infty} (d_{k+1} - d_0)$ or of the sequence (d_n) (see Hardy [11], p. 80, or Zeller and Beekmann [30], p. 134).

Lemma 4 Let the sequence $(c_n)_{n \in \mathbb{N}}$ of real numbers satisfy

$$\sum_{n=1}^{\infty} n^{-3/2} c_n^2 < \infty.$$
 (20)

If

$$\int_{0}^{\infty} t^{-1} |e^{-t} \sum_{k=1}^{\infty} \frac{t^{k}}{k!} c_{k}|^{2} dt < \infty$$
(21)

or only

$$\frac{1}{\lambda} \int_{0}^{\lambda} e^{-t} \sum_{k=1}^{\infty} \frac{t^{k}}{k!} c_{k} dt \to 0 \quad (\lambda \to \infty),$$
(22)

then

$$\frac{1}{n}\sum_{i=1}^{n}c_{i} \to 0 \quad (n \to \infty), \tag{23}$$

Remark 6 As in the proof of Theorem 3 given in section 3, one can conclude from Proposition 1 that the condition

$$|\Gamma(i,j)| \le c \frac{i^{\alpha} + j^{\alpha}}{1 + |i-j|^{\beta}}, \ i, j \in \mathbb{N},$$
(24)

for some $c \in \mathbb{R}_+$, $0 < 2\alpha < \beta < 1$ is sufficient for (11) and thus (1). This condition for the strong law of large numbers was established by Cramér and Leadbetter [5], p. 94. It seems to be an open problem whether the condition can be relaxed to $0 < \alpha < \beta < 1$. Ninness [24] in his argument does not verify completely the conditions of Lemma 1 there. The relaxed condition yields

$$\sum_{i,j\in\{1,\dots,n\}} |\Gamma(i,j)| \le c^* n^{2+(\alpha-\beta)}, n \in \mathbb{N}$$

for some $c^* \in \mathbb{R}_+$ (see Ninness [24], p. 219, and (3), via (4)). Under the additional assumption (2), now (1) is obtained by Theorem 1.

Remark 7 In a straightforward way Theorems 1 and 3 and their proofs can be transferred to the case of stochastic processes $\{X(t); t \in \mathbb{R}_+(\text{or } \mathbb{R})\}$ that are continuous in squared mean, where sums are replaced by integrals. Especially one obtains a generalization of Loève's [21] continuous time version of his result mentioned in Remark 1 a.

3 PROOFS

Assumption (6) in Lemma 1 implies

$$\frac{1}{n^3} \sum_{k=1}^n (c_1 + \ldots + c_k)^2 \to 0$$

by the Kronecker lemma and thus assumption (7) by the Cauchy-Schwarz inequality. (7) is equivalent to

$$\frac{1}{\binom{n+1}{2}} \sum_{k=1}^{n} (n+1-k)c_k \to 0,$$
(25)

i.e., C_2 -summability of the sequence (c_n) to 0, while the assertion (8) means C_1 -summability of (c_n) to 0 (see Hardy [11], p. 96, p. 7, or Zeller and Beekmann [30] p. 100, p. 104). With

$$(1-s)\sum_{k=1}^{\infty} c_k s^{k-1} \to 0 \ (s\uparrow 1),$$
 (26)

i.e., Abel summability of (c_n) to 0, one has $(8) \Longrightarrow (25) \Longrightarrow (26)$ for general sequences (c_n) of real numbers (see Hardy [11], p. 7, Theorems 43 and 55, or Zeller and Beekmann [30], p. 110, 53 I and 55 II). For one-side bounded sequences (c_n) also the converse holds. In this situation, in contrast to the implication $(26) \Longrightarrow (8)$ (see Hardy [11], Theorem 97, or Zeller and Beekmann [30], 55 IV, or Feller [8], XIII.5, Theorem 5) the implication (25) (or (7)) \Longrightarrow (8) can be proved in an elementary way. To make the paper self-contained, we give the proof of Lemma 1, i.e., of $(7) \Longrightarrow (8)$ for (c_n) bounded from below, according to Mordell [23] and Boas [1], proof of Theorem 4b there (in these papers with $\phi(n) = n^2$, $\psi(n) = -1$ and 1, respectively), compare also Zeller and Beekmann [30], p. 117, for further references.

Proof of Lemma 1. It suffices to conclude (8) from (7) under the assumption $c_n \ge -c$ $(n \in \mathbb{N}), c \in \mathbb{R}_+$. Set

$$t_n := \sum_{k=1}^n c_k, \ w_n := \sum_{k=1}^n t_k, \ n \in \mathbb{N}.$$

For $k \in \{1, \ldots, n\}$ one has

$$\left. \begin{array}{cc} \sum\limits_{j=1}^{k} & (t_{n+j} - t_n) \\ \sum\limits_{j=0}^{k-1} & (t_n - t_{n-j}) \end{array} \right\} \ge -c \sum\limits_{j=1}^{k} j \ge -k^2 c,$$

thus

$$w_{n\pm k} \geq w_n \pm kt_n - k^2 c,$$

$$\pm \frac{t_n}{n} \leq \frac{w_{n\pm k} - w_n}{kn} + \frac{k}{n} c.$$
(27)

One notices

$$\sigma_n := \max\{|w_j|; j = 1, \dots, 2n\} = o(n^2) \text{ (by (7))}, \\ \delta_n := 1 + \lfloor \sqrt{\sigma_n} \rfloor = o(n),$$

($\lfloor \ \ \rfloor$ denoting integer part). Now from (27) with $k = \delta_n$ (< n for n sufficiently large), one obtains

$$\limsup(\pm \frac{t_n}{n}) \le \lim\left(\frac{2\sigma_n}{n\delta_n} + \frac{\delta_n}{n}c\right) = 0,$$
(8).

i.e. $t_n/n \to 0$, i.e., (8).

Proof of Theorem 1. Using the monotone convergence theorem in series form, we obtain

$$E\sum_{n=1}^{\infty} \frac{|\sum_{j=1}^{n} (X_j - EX_j)|^2}{n^3} = \sum_{n=1}^{\infty} \frac{E|\sum_{j=1}^{n} (X_j - EX_j)|^2}{n^3}$$
$$= \sum_{n=1}^{\infty} \frac{\operatorname{Var}(X_1 + \dots + X_n)}{n^3}$$
$$< \infty \text{ (by (2.2)).}$$

Thus

$$\sum_{n=1}^{\infty} \frac{(\sum_{j=1}^{n} (X_j - EX_j))^2}{n^3} < \infty \text{ a.s.}$$

By (2) the sequence $(X_n - EX_n)$ is bounded from below a.s. Now Lemma 1 yields the assertion.

To make the paper self-contained we repeat the following well-known proof.

Proof of Lemma 2.

$$\sum_{n=1}^{\infty} \frac{E(X^{[n]^2})}{n^2} = \sum_{n=1}^{\infty} \sum_{i=1}^{n} \frac{1}{n^2} \int_{(i-1,i]} t^2 P_X(dt)$$
$$= \sum_{i=1}^{\infty} \int_{(i-1,i]} t^2 P_X(dt) \sum_{n=i}^{\infty} \frac{1}{n^2}$$
$$\leq \sum_{i=1}^{\infty} \frac{2}{i} \int_{(i-1,i]} t^2 P_X(dt)$$
$$\leq 2EX < \infty.$$

Proof of Theorem 2. Assume $X_n \ge 0$ without loss of generality. The first step is well-known. Set

$$X_n^* := X_n^{[n]} := X_n I_{[X_n \le n]}, \ n \in \mathbb{N}.$$

Because of

$$\sum_{i=1}^{\infty} P\left[X_i \neq X_i^*\right] = \sum_{i=1}^{\infty} P\left[X_1 > i\right] \le EX_1 < \infty,$$

a.s. $X_i = X_i^*$ from some index on (by the Borel-Cantelli lemma; Loève [21], sections 16,17). Therefore and because of

$$EX_n^* = EX_1^{[n]} \to EX_1 < \infty$$

(by the monotone convergence theorem), it suffices to show

$$\frac{1}{n} \sum_{i=1}^{n} (X_i^* - EX_i^*) \to 0 \text{ a.s.}$$
(28)

In the second step we notice

$$X_n^* \ge 0, \ EX_n^* \le EX_1 < \infty,$$

further,

$$\sum_{n=1}^{\infty} \frac{\operatorname{Var}(X_{1}^{*} + \ldots + X_{n}^{*})}{n^{3}}$$

$$\leq \sum_{n=1}^{\infty} \frac{\operatorname{Var}(X_{1}^{*}) + \ldots + \operatorname{Var}(X_{n}^{*})}{n^{3}} + 2\sum_{n=1}^{\infty} \frac{1}{n^{3}} \sum_{j=2}^{n} \sum_{i=1}^{j-1} \kappa(j-i,j) \sqrt{\operatorname{Var}(X_{j-i})} \sqrt{\operatorname{Var}(X_{j})}$$

$$\leq \sum_{n=1}^{\infty} \frac{EX_{n}^{*2}}{n^{2}} \left(1 + 2 \sup_{j \in \{2,3,\ldots\}} \sum_{i=1}^{j-1} \kappa(j-i,j) \right)$$

$$< \infty \text{ (by Lemma 2 and (10)).}$$

Thus (28) follows by Theorem 1.

Proof of Lemma 3. We use the abbreviation

$$v_{n,k} := e^{-n} \frac{n^k}{k!} \ (n \in \mathbb{N}, k \in \mathbb{N}_0).$$

By Hölder's inequality, for arbitrary $M \in \mathbb{N}$ we obtain

$$\sum_{k=M+1}^{\infty} v_{n,k} |d_k - d_n| \le \left(\sum_{k=M+1}^{\infty} v_{n,k} |d_k - d_n|^4\right)^{1/4}$$

noticing

$$\sum_{k=0}^{\infty} v_{n,k} = 1.$$
 (29)

We now modify an argument of Jurkat [16], p. 280. We set

$$u_j := d_j - d_{j-1}, \ j \in \mathbb{N}_0$$

and obtain

$$\begin{aligned} |d_{l} - d_{i}|^{2} &= \left| \sum_{j=i+1}^{l} u_{j} \right|^{2} \\ &\leq \sum_{j=i+1}^{l} j^{-1/2} \sum_{j=i+1}^{l} j^{1/2} u_{j}^{2} \\ &\quad \text{(by the Cauchy-Schwarz inequality)} \\ &\leq 2 \frac{l-i}{l^{1/2} + i^{1/2}} \sum_{j=i+1}^{l} j^{1/2} u_{j}^{2} \quad (1 \leq i < l), \end{aligned}$$
(30)

thus for $n > M (\in \mathbb{N})$

$$\left(\sum_{k=M+1}^{\infty} v_{n,k} |d_k - d_n|\right)^4 \\
\leq 4 \sum_{k=M+1}^{\infty} v_{n,k} \frac{|k-n|^2}{n} \left(\sum_{j=M+1}^{\infty} j^{1/2} u_j^2\right)^2 \\
\leq 4 \left(\sum_{j=M+1}^{\infty} j^{1/2} u_j^2\right)^2,$$
(31)

the latter because of

$$\sum_{k=0}^{\infty} v_{n,k}(k-n)^2 = n, \ n \in \mathbb{N}.$$

For arbitrary $\varepsilon > 0$, by (31) and (17),

$$\sum_{k=M+1}^{\infty} v_{n,k} |d_k - d_n| < \varepsilon$$

if M is sufficiently large and n > M. Further for each $k \in \mathbb{N}_0$

$$v_{n,k} \to 0, \qquad v_{n,k} |d_n| \to 0 \ (n \to \infty),$$

because (17) yields $d_n - d_{n-1} = o(n^{-1/4})$, and thus $d_n = o(n^{3/4})$. Therefore

$$\sum_{k=0}^{\infty} v_{n,k} |d_k - d_n| \to 0 \quad (n \to \infty).$$

This together with (18) and (29) yields (19).

Lemma 3 can also be obtained by a classic and deep Tauberian theorem of R. Schmidt [27] (see Hardy [11], p. 225, p. 312, and Zeller and Beekmann [30], 66 X). It states that (19) is implied by (18) together with

$$\liminf(d_n - d_m) \ge 0 \tag{32}$$

where

$$m \to \infty, n > m, \frac{n-m}{m^{1/2}} \to 0.$$

(32) is weaker than (17), because of (30).

Proof of Lemma 4. Set $c_0 := 0$ and

$$d_j := \frac{1}{j+1} \sum_{k=0}^{j} c_k, \ j \in \mathbb{N}_0.$$
(33)

Then

$$\sum_{n=1}^{\infty} n^{1/2} (d_n - d_{n-1})^2$$

$$\leq 2 \sum_{n=1}^{\infty} n^{-3/2} c_n^2 + 2 \sum_{n=1}^{\infty} n^{-7/2} \left(\sum_{k=1}^n c_k \right)^2$$

$$\leq 2 \sum_{n=1}^{\infty} n^{-3/2} c_n^2 + 2 \sum_{n=1}^{\infty} n^{-5/2} \sum_{k=1}^n c_k^2.$$

Changing the order of summation, we obtain (17) by (20). From (21) we obtain

$$\frac{1}{\lambda} \int_{0}^{\lambda} |e^{-t} \sum_{k=1}^{\infty} \frac{t^{k}}{k!} c_{k}|^{2} dt \to 0 \quad (\lambda \to \infty)$$

via partial integration as in the proof of the classical Kronecker lemma (see, e.g., Loève [21], section 17), and then (22) by the Cauchy-Schwarz inequality. Thus it suffices to assume (22). This assumption means that the integral Cesàro limit of the Borel transformation

$$\lambda \to e^{-\lambda} \sum_{k=0}^{\infty} \frac{\lambda^k}{k!} c_k, \ \lambda \ge 0$$

of $(c_k)_{k \in \mathbb{N}_0}$ is 0. As is well known, Cesàro transformation and Borel transformation are commutative for sequences whose Borel transformation exists (see Zeller and Beekmann [30], p. 139). In fact

$$\int_{0}^{\lambda} \frac{1}{k!} t^{k} e^{-t} dt = e^{-\lambda} \sum_{j=k+1}^{\infty} \frac{1}{j!} \lambda^{j}, \ k \in \mathbb{N}_{0}, \lambda > 0$$

(proof by differentiation), thus for the considered sequence $(c_k)_{k \in \mathbb{N}_0}$

$$\frac{1}{\lambda} \int_{0}^{\lambda} \sum_{k=0}^{\infty} \frac{1}{k!} t^{k} e^{-t} c_{k} dt$$

$$= e^{-\lambda} \sum_{k=0}^{\infty} \sum_{j=k+1}^{\infty} \frac{1}{j!} \lambda^{j-1} c_{k}$$

$$= e^{-\lambda} \sum_{j=1}^{\infty} \left(\sum_{k=0}^{j-1} c_{k} \right) \frac{1}{j!} \lambda^{j-1}$$

$$= e^{-\lambda} \sum_{j=0}^{\infty} \frac{1}{j!} \lambda^{j} \left(\frac{1}{j+1} \sum_{k=0}^{j} c_{k} \right), \ \lambda > 0.$$

Therefore (22) and (18) with (33) are equivalent.

Now we use Lemma 3. With (33), from (18) and (17) we obtain (19), i.e., the assertion (23). \Box

Proof of Proposition 1. We assume $EX_n = 0, n \in \mathbb{N}$, without loss of generality. Condition (12) together with $\Gamma(j, j) = EX_j^2$ immediately yields

$$\sum n^{-3/2} E X_n^2 < \infty \tag{34}$$

and thus

$$\sum n^{-3/2} X_n^2 < \infty \text{ a.s.}$$
(35)

We shall show

$$I := \int_{0}^{\infty} t^{-1} E |e^{-t} \sum_{k=1}^{\infty} \frac{t^{k}}{k!} X_{k}|^{2} dt < \infty$$
(36)

and thus

$$\int_{0}^{\infty} t^{-1} |e^{-t} \sum_{k=1}^{\infty} \frac{t^{k}}{k!} X_{k}|^{2} dt < \infty \text{ a.s.}$$
(37)

Apparently

$$I = \int_{0}^{\infty} t^{-1} e^{-2t} \sum_{j,l \in \mathbb{N}} \frac{t^{j}}{j!} \frac{t^{l}}{l!} \Gamma(l,j) dt$$

$$\leq 2 \int_{0}^{\infty} t^{-1} e^{-2t} \sum_{k=0}^{\infty} \sum_{j=1}^{\infty} \frac{t^{2j}}{j!} \frac{t^{k}}{(j+k)!} |\Gamma(j+k,j)| dt$$

$$= 2 \sum_{k=0}^{\infty} s_{k}$$

with

$$s_k = \sum_{j=1}^{\infty} (2j+k-1)! \ 2^{-(2j+k)} \frac{1}{j!} \frac{1}{(j+k)!} |\Gamma(j+k,j)|, \ k \in \mathbb{N}_0,$$

via the Gamma function. Now, by Stirling's formula, we obtain

$$s_k \leq c_1 \sum_{j=1}^{\infty} \frac{(2j+k)^{2j+k}}{j^j (j+k)^{j+k} \sqrt{j} (j+k)} 2^{-(2j+k)} |\Gamma(j+k,j)|$$

= $c_1 \sum_{j=1}^{\infty} \frac{(1+k/(2j))^{2j+k}}{(1+k/j)^{j+k} \sqrt{j} (j+k)} |\Gamma(j+k,j)|$

with suitable $c_1 > 0$. By (34) we have

$$s_0 < \infty \tag{38}$$

We notice for $0 < s < \infty$

$$0 > h(s) := \left(\frac{2}{s} + 1\right) \ln\left(1 + \frac{s}{2}\right) - \left(\frac{1}{s} + 1\right) \ln\left(1 + s\right)$$
$$= \begin{cases} -(s/4) + o(s) & (s \to 0) \\ -\ln 2 + o(1) & (s \to \infty), \end{cases}$$

thus, for suitable c > 0,

$$\frac{(1+s/2)^{(2/s)+1}}{(1+s)^{(1/s)+1}} \le \begin{cases} e^{-cs} & , & 0 < s \le 1\\ e^{-c} & , & s \ge 1. \end{cases}$$

Noticing convexity of h we can choose c = 1/6. Then, with s = k/j $(k \in \mathbb{N}, j \in \mathbb{N})$ we obtain

$$s_k \leq c_1 \left[e^{-ck} \sum_{j=1}^k j^{-3/2} |\Gamma(j+k,j)| + \sum_{j=k+1}^\infty e^{-ck^2/j} j^{-\frac{3}{2}} |\Gamma(j+k,j)| \right]$$

= $s'_k + s''_k$.

From (34) we obtain

$$\begin{aligned} |\Gamma(j+k,j)| &\leq |\Gamma(j+k,j+k)|^{1/2} |\Gamma(j,j)|^{1/2} \\ &= O((j+k)^{3/2}) \end{aligned}$$

and thus

$$\sum_{k=1}^{\infty} s_k' < \infty. \tag{39}$$

By (12) we have

$$\sum_{k=1}^{\infty} s_k'' < \infty \tag{40}$$

(38), (39), and (40) yield (36). Finally from (35) together with (37) we obtain (1) by Lemma 4. $\hfill \Box$

Proof of Theorem 3. We use Proposition 1. It is enough to show

$$\sum_{k=1}^{\infty} \int_{k}^{\infty} e^{-k^2/(6t)} t^{-3/2} dt \ r(k) < \infty$$

(because of (13) and piecewise monotonicity of the integrand). But this follows from

$$\int_{k}^{\infty} e^{-k^{2}/(6t)} t^{-3/2} dt = \int_{0}^{k} e^{-v/6} v^{-1/2} dv \ k^{-1} \text{ and } (14).$$

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2004/001 *Walk, H.*: Strong Laws of Large Numbers by Elementary Tauberian Arguments.